

Empirical Bayes learning from co-data in high-dimensional prediction settings

Mark van de Wiel^{1,2}

¹Dep of Epidemiology and Biostatistics, VU University medical center (VUmc)

²Dep of Mathematics, VU University, Amsterdam, The Netherlands

Contributions by: **Putri Novianti** (VUmc), **Magnus Münch** (Leiden, VUmc)

Our group: www.bigstatistics.nl

Setting

- **Prediction or Classification**
- **Primary data**
 - ▶ Variables $i = 1, \dots, p$; Individuals $j = 1, \dots, n$; $p > n$
 - ▶ Focus on binary response Y_j (e.g. case vs control)
 - ▶ Measurements $\mathbf{X}_j = (X_{1j}, \dots, X_{pj})$
 - ▶ Goal: find f such that $Y_j \approx f(\mathbf{X}_j)$
 - ▶ f : *logistic regression*, random forest, spike-and-slab, etc.
 - ▶ Some form of regularization required
- **Focus**
 - ▶ Differential regularization based on prior information:
Co-data

Co-data

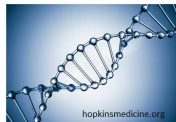
Definition Co-data: any information on the *variables* not using the response labels of the primary data

Co-data

Definition Co-data: any information on the *variables* not using the response labels of the primary data

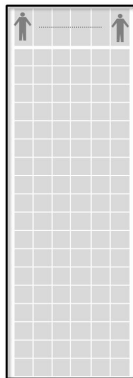
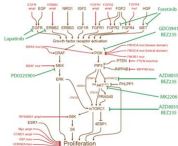


Databases



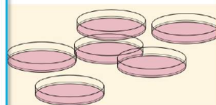
Related bio-molecules

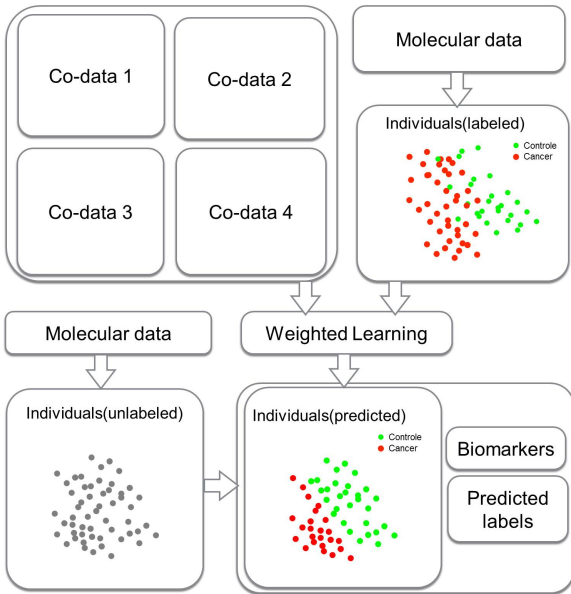
Pathways



Primary Data

Cell lines





Use of co-data

Groups: Co-data determine G prior groups of variables

Idea: Use different penalty weights $\lambda_1, \dots, \lambda_G$ across G co-data-based groups. E.g. in ridge:

$$\operatorname{argmax}_{\beta} \mathcal{L}(\mathbf{Y}; \beta) - \sum_{g=1}^G \lambda_g \|\beta_g\|_2$$

Use of co-data

Groups: Co-data determine G prior groups of variables

Idea: Use different penalty weights $\lambda_1, \dots, \lambda_G$ across G co-data-based groups. E.g. in ridge:

$$\operatorname{argmax}_{\beta} \mathcal{L}(\mathbf{Y}; \beta) - \sum_{g=1}^G \lambda_g \|\beta_g\|_2$$

Challenge: Estimation of hyperparameters λ_g

CV not attractive

Empirical Bayes (EB)

Definition*: EB estimates the prior from data

→ Parametric form: estimate prior parameters

→ Penalized regression: estimate penalty parameters; via link with prior

*Excellent discussions: Carlin & Louis (2000), Efron (2010), Van Houwelingen (2014)

Empirical Bayes (EB)

Definition*: EB estimates the prior from data

→ Parametric form: estimate prior parameters

→ Penalized regression: estimate penalty parameters; via link with prior

Why Empirical Bayes (EB)?

- EB estimators tend to improve for increasing p

*Excellent discussions: Carlin & Louis (2000), Efron (2010), Van Houwelingen (2014)

Empirical Bayes (EB)

Definition*: EB estimates the prior from data

→ Parametric form: estimate prior parameters

→ Penalized regression: estimate penalty parameters; via link with prior

Why Empirical Bayes (EB)?

- EB estimators tend to improve for increasing p
- EB fits well with allowing for prior information: can improve predictions

*Excellent discussions: Carlin & Louis (2000), Efron (2010), Van Houwelingen (2014)

Empirical Bayes (EB)

Definition*: EB estimates the prior from data

→ Parametric form: estimate prior parameters

→ Penalized regression: estimate penalty parameters; via link with prior

Why Empirical Bayes (EB)?

- EB estimators tend to improve for increasing p
- EB fits well with allowing for prior information: can improve predictions
- Computationally nicer than Full Bayes and CV

*Excellent discussions: Carlin & Louis (2000), Efron (2010), Van Houwelingen (2014)

Formal EB: Maximum marginal Likelihood

$\beta = (\beta_1, \dots, \beta_p)$. Prior(s): $\pi_{\alpha}(\beta)$, $\alpha = (\alpha_1, \dots, \alpha_K)$

Marginal likelihood maximization:

$$\hat{\alpha} = \operatorname{argmax}_{\alpha} \operatorname{ML}(\alpha), \text{ with } \operatorname{ML}(\alpha) = \int_{\beta} \mathcal{L}(\mathbf{Y}; \beta) \pi_{\alpha}(\beta) d\beta,$$

Formal EB: Maximum marginal Likelihood

$\beta = (\beta_1, \dots, \beta_p)$. Prior(s): $\pi_{\alpha}(\beta)$, $\alpha = (\alpha_1, \dots, \alpha_K)$

Marginal likelihood maximization:

$$\hat{\alpha} = \operatorname{argmax}_{\alpha} \operatorname{ML}(\alpha), \text{ with } \operatorname{ML}(\alpha) = \int_{\beta} \mathcal{L}(\mathbf{Y}; \beta) \pi_{\alpha}(\beta) d\beta,$$

High-dimensional integral \rightarrow optimization hard

High-dimensional integral

Solutions:

- Laplace approximation (Shun & McCullagh, 1995)
- EM on Gibbs samples (Casella, 2001). Conceptually easy, but computationally very intensive.
- EM on Variational Bayes approximation (Bernardo et al., 2003). Fast, but dedicated approximations[†].

[†]Work in progress for elastic net and spike-and-slab

High-dimensional integral

Solutions:

- Laplace approximation (Shun & McCullagh, 1995)
- EM on Gibbs samples (Casella, 2001). Conceptually easy, but computationally very intensive.
- EM on Variational Bayes approximation (Bernardo et al., 2003). Fast, but dedicated approximations[†].
- **Or** resort to alternative EB approach

[†]Work in progress for elastic net and spike-and-slab

Back to the ridge example

Empirical Bayes (EB) estimation of λ_g explores

$$\operatorname{argmax}_{\beta} \mathcal{L}(\mathbf{Y}; \beta) - \sum_{g=1}^G \lambda_g \|\beta_g\|_2 = \beta_{\text{MAP}},$$

when

$$j \in \text{Group } g : \beta_j \sim \mathbf{N}(0, \tau_g^2), \tau_g^{-2} \propto \lambda_g$$

Back to the ridge example

Empirical Bayes (EB) estimation of λ_g explores

$$\operatorname{argmax}_{\beta} \mathcal{L}(\mathbf{Y}; \beta) - \sum_{g=1}^G \lambda_g \|\beta_g\|_2 = \beta_{\text{MAP}},$$

when

$$j \in \text{Group } g : \beta_j \sim \mathbf{N}(0, \tau_g^2), \tau_g^{-2} \propto \lambda_g$$

→ EB estimate of τ_g^2 renders estimate of λ_g .

EB for group-regularized ridge[‡]

Aim: $\hat{\tau}_g^2$ for group-regularized ridge: $\beta_i \sim N(0, \tau_g^2), i \in \mathcal{G}_g$

[‡]Details: Van de Wiel et al., *Stat Med*, 2016

EB for group-regularized ridge[‡]

Aim: $\hat{\tau}_g^2$ for group-regularized ridge: $\beta_i \sim N(0, \tau_g^2), i \in \mathcal{G}_g$

Initial: $\hat{\beta}_i = \hat{\beta}_i^{\lambda_0}$. Moment equations $g = 1, \dots, G$. $G = 2$:

$$\frac{1}{p_1} \sum_{i \in \mathcal{G}_1} \hat{\beta}_i^2 \approx \frac{1}{p_1} \sum_{i \in \mathcal{G}_1} E_{\beta} \left[E[\hat{\beta}_i^2(\mathbf{Y}) | \beta] \right] := h_1(\tau_1^2, \tau_2^2)$$

$$\frac{1}{p_2} \sum_{i \in \mathcal{G}_2} \hat{\beta}_i^2 \approx \frac{1}{p_2} \sum_{i \in \mathcal{G}_2} E_{\beta} \left[E[\hat{\beta}_i^2(\mathbf{Y}) | \beta] \right] := h_2(\tau_1^2, \tau_2^2),$$

[‡]Details: Van de Wiel et al., *Stat Med*, 2016

EB for group-regularized ridge[‡]

Aim: $\hat{\tau}_g^2$ for group-regularized ridge: $\beta_i \sim N(0, \tau_g^2), i \in \mathcal{G}_g$

Initial: $\hat{\beta}_i = \hat{\beta}_i^{\lambda_0}$. Moment equations $g = 1, \dots, G$. $G = 2$:

$$\frac{1}{p_1} \sum_{i \in \mathcal{G}_1} \hat{\beta}_i^2 \approx \frac{1}{p_1} \sum_{i \in \mathcal{G}_1} E_{\beta} \left[E[\hat{\beta}_i^2(\mathbf{Y}) | \beta] \right] := h_1(\tau_1^2, \tau_2^2)$$

$$\frac{1}{p_2} \sum_{i \in \mathcal{G}_2} \hat{\beta}_i^2 \approx \frac{1}{p_2} \sum_{i \in \mathcal{G}_2} E_{\beta} \left[E[\hat{\beta}_i^2(\mathbf{Y}) | \beta] \right] := h_2(\tau_1^2, \tau_2^2),$$

In general: System of G linear equations $\mathbf{b}_{\text{data}} = \mathbf{A}t$

Solution: $t = (\hat{\tau}_1^2, \dots, \hat{\tau}_G^2)$.

[‡]Details: Van de Wiel et al., *Stat Med*, 2016

Extension: Stability §

- Some co-data render *many* groups: e.g. pathways
- G large: system $\mathbf{b}_{\text{data}} = \mathbf{A}\mathbf{t}$ becomes unstable
- Need to stabilize solution

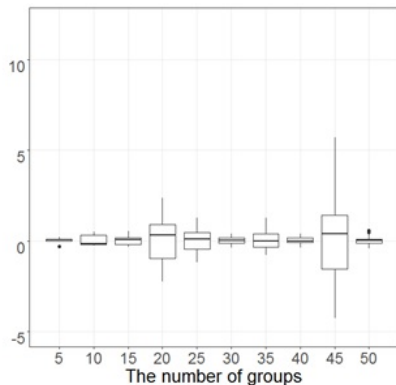
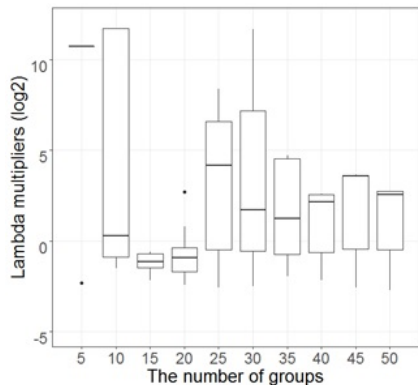
Solutions

1. Enforce monotony when grouping based on continuous co-data (e.g. external p-values)
2. Shrink A to a stable target T : $\tilde{A}_q = qA + (1 - q)T$.

§Details: Novianti et al., *Bioinformatics*, 2017

Effect of shrinkage of A

Real data, *random* groups of variables



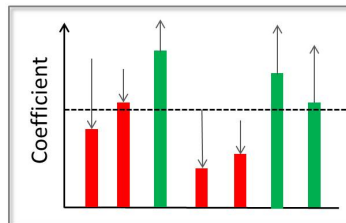
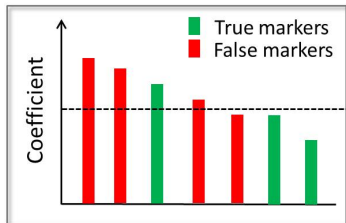
Left: No Shrinkage; Right: Shrinkage

Variable selection

Why can co-data help?

No Co-data

With Co-data



Variable selection

Current solution:

1. Estimate group penalties from ridge regression, possibly for multiple groupings
2. Select k variables by introducing non-grouped L_1 penalty (i.e. thresholding)
3. Refit the model using the selected variables and their respective L_2 penalties

Variable selection

Current solution:

1. Estimate group penalties from ridge regression, possibly for multiple groupings
2. Select k variables by introducing non-grouped L_1 penalty (i.e. thresholding)
3. Refit the model using the selected variables and their respective L_2 penalties

“**Bet on sparsity**”: yes, but *after* penalty weighting

R-package `GRRidge`, Github + Bioconductor:

- Logistic, linear and survival
- Auxiliary functions for co-data processing (from TCGA etc.)
- Allows unpenalized covariates
- Built-in CV for comparison with ridge & lasso

¶Details: Novianti et al., *Bioinformatics*, 2017

R-package `GRRidge`, Github + Bioconductor:

- Logistic, linear and survival
- Auxiliary functions for co-data processing (from TCGA etc.)
- Allows unpenalized covariates
- Built-in CV for comparison with ridge & lasso

Comparison (one grouping only): Sparse group lasso, SGL (Simon et al., *J Comp Graph Stat*, 2013).

[¶]Details: Novianti et al., *Bioinformatics*, 2017

Example: Diagnostics for cervical cancer

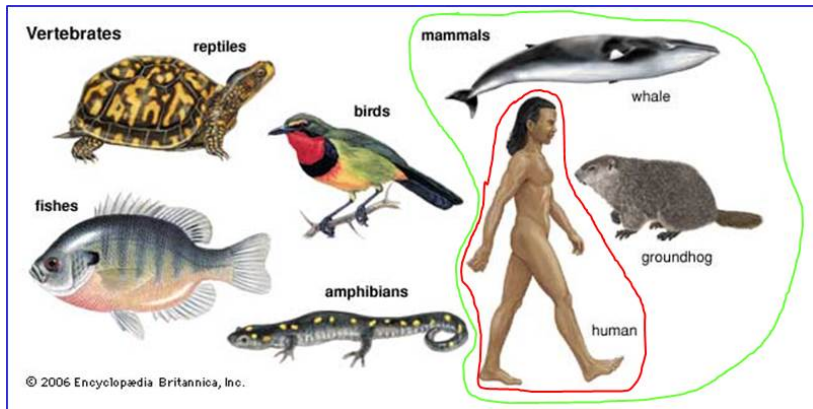
Goal: Select markers for classifying Normal vs CIN3
→ final goal is a cheap PCR assay

Data:

- microRNA sequencing data
- $n = 56$: 32 Normal, 24 CIN3
- $p = 772$ (after filtering lowly abundant ones).
- Sqrt-transformed
- Standardized

Co-data 1: Conservation status

1. Non-conserved, **human** only (552)
2. Conserved across **mammals** (72)
3. Broadly conserved, across most **vertebrates** (148)



Co-data 2: Standard deviation

- Current practice: standardize variable j by sd: s_j
 - effective penalty λs_j^2 (Zwiener et al, 2014)
 - too large advantage for small s_j 's?

Co-data 2: Standard deviation

- Current practice: standardize variable j by sd: s_j
 - effective penalty λs_j^2 (Zwiener et al, 2014)
 - too large advantage for small s_j 's?
- Our solution:
 1. Standardize by s_j
 2. G groups of variables with decreasing s_j
 3. Effective penalty $j \in \mathcal{G}_g$: $\lambda_j = \hat{\tau}_g^{-2} \lambda s_j^2$

Co-data 2: Standard deviation

- Current practice: standardize variable j by sd: s_j
 - effective penalty λs_j^2 (Zwiener et al, 2014)
 - too large advantage for small s_j 's?
- Our solution:
 1. Standardize by s_j
 2. G groups of variables with decreasing s_j
 3. Effective penalty $j \in \mathcal{G}_g$: $\lambda_j = \hat{\tau}_g^{-2} \lambda s_j^2$
- Allows a more non-parametric link between s_j and λ_j

Co-data results

For $j \in \mathcal{G}_g$, penalty factor: $\lambda'_g \propto \tau_g^{-2}$

Co-data results

For $j \in \mathcal{G}_g$, penalty factor: $\lambda'_g \propto \tau_g^{-2}$

Conservation status:

1. Non-conserved (552): $\lambda'_1 = 1.84$
2. Conserved across mammals (72): $\lambda'_2 = 0.61$
3. Broadly conserved across vertebrates (148): $\lambda'_3 = 0.30$

Co-data results

For $j \in \mathcal{G}_g$, penalty factor: $\lambda'_g \propto \tau_g^{-2}$

Conservation status:

1. Non-conserved (552): $\lambda'_1 = 1.84$
2. Conserved across mammals (72): $\lambda'_2 = 0.61$
3. Broadly conserved across vertebrates (148): $\lambda'_3 = 0.30$

Standard deviation:

Range from $\lambda'_1 = 0.56$ (large s.d.) to $\lambda'_{10} = 1.80$ (small s.d.)

→ Indeed, partly ‘undoes’ the effect of standardization (for $j \in \mathcal{G}_g$: $\lambda_j \propto \lambda'_g s_j^2$).

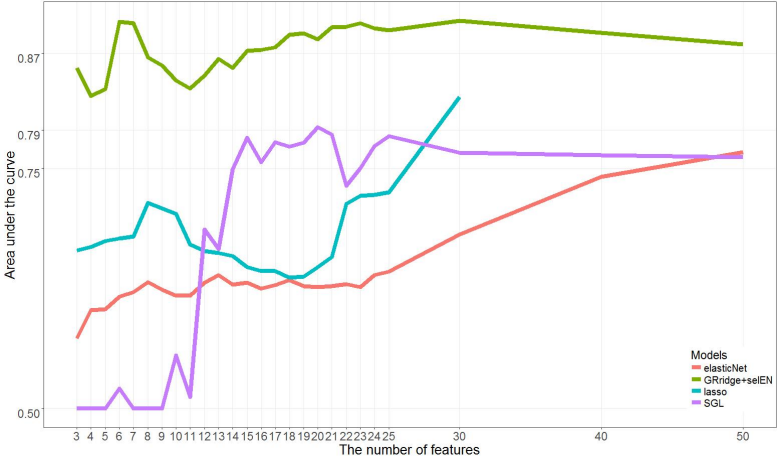
Clinician:

“That’s all nice, but does the predictive accuracy improve?”

“Do I get the good biomarkers?”

Performance under variable selection

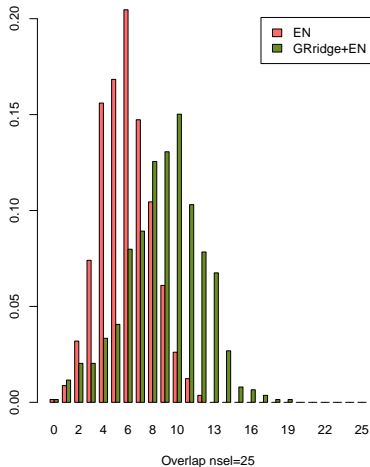
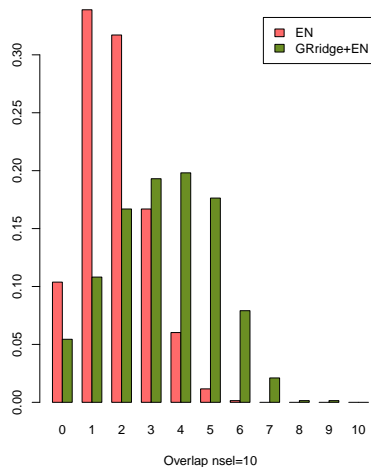
AUC assessed by LOOCV



GRridge + EN, Sparse group-lasso, Lasso, Elastic Net

Stability of selection

50 re-sampled versions of data set. Overlap in selected variables between pairs of re-samples



Other applications, extensions

Hybrid Bayes - Empirical Bayes

- $\lambda_g = \lambda'_g \lambda$. λ'_g : EB; Common λ : Full Bayes (prior)

|| *Ann Appl Stat*, 2016; *Biom J*, 2017

** arXiv, 2017

Other applications, extensions

Hybrid Bayes - Empirical Bayes

- $\lambda_g = \lambda'_g \lambda$. λ'_g : EB; Common λ : Full Bayes (prior)

Networks (Gwenaël Leday, Gino Kpogbezan et al. ^{||})

- Bayesian SEM: Variational Bayes + EB + prior network

^{||} *Ann Appl Stat*, 2016; *Biom J*, 2017

^{**} arXiv, 2017

Other applications, extensions

Hybrid Bayes - Empirical Bayes

- $\lambda_g = \lambda'_g \lambda$. λ'_g : EB; Common λ : Full Bayes (prior)

Networks (Gwenaël Leday, Gino Kpogbezan et al. ^{||})

- Bayesian SEM: Variational Bayes + EB + prior network

Random Forest (Dennis te Beest ^{**})

- Co-data moderated Random Forest

^{||} *Ann Appl Stat*, 2016; *Biom J*, 2017

^{**} arXiv, 2017

Take home

Empirical Bayes...

... is a versatile technique to learn

- 1. from a lot...(many variables)**
- 2. ...and a lot more (co-data)**

Acknowledgements



Magnus Münch
(Bayes EN)

Dennis te Beest
(RF)

Gino Kpogbezan
(Networks)

Wessel "Ridge"
van Wieringen

Putri Novianti (GRridge)

+ miRNAseq data: Barbara Snoek, Saskia Wilting, Renske Steenbergen (VUmc)

+ Grant: NWO ZONMW-TOP

Details

Method: Van de Wiel MA, Lien TG, Verlaat W, Van Wieringen WN, Wilting SM (2016). Better prediction by use of co-data: Adaptive group-regularized ridge regression. *Stat Med.*, **35**, 368-381.

Software: Novianti PW, Snoek B, Wilting SM, van de Wiel MA (2017). Better diagnostic signatures from RNAseq data through use of auxiliary co-data. *Bioinformatics*, **33**, 1572-1574.

QUESTIONS?^{††}